# Political Economics II Spring 2016 

# Part III: Political Institutions and Economic Policy 

Torsten Persson, IIES

## Introduction

Effects of political institutions - "constitutional reform"?
what happens to economic policy?
traditionally, neglected between economics and political science
Research in last ten-fifteen years
new theory developed (applying tools you have seen already) and used as stepping stone for empirical analysis

Goal of Lectures 7-8
give selective introduction to research results and methods underlying theoretical models and predictions (Lecture 7) empirical measurement, strategy and results (Lecture 8)

Reforms of which constitutional features?
two reasonably fundamental features, where theory exists electoral rules: "majoritarian" vs. "proportional" (and finer detail) forms of government: "presidential" vs. "parliamentary"

Which policies?
again, follow theory: fiscal policy (rent extraction)
large variation in observed outcomes, even among similar countries
Domain of application?
rules and policy outcomes at national level
parallel literature about sub-national institutions and policy

## Background and Theoretical Predictions

today, we begin with some theory and its predictions
Agenda
A. A public-finance problem
B. Electoral rules and political competition
C. Forms of government and legislative bargaining
D. Summary of predictions

## A. A public-finance problem

Groups of voters
homogenous (continuum) within $J=1,2,3$, equal size, $N^{J}=1$
Private preferences over economic outcomes

$$
\begin{equation*}
w^{J}=c^{J}+H(g)=y+(1-\tau)+f^{J}+H(g) \tag{1}
\end{equation*}
$$

policies both targeted $\left(f^{J}\right)$ and non-targeted $(\tau, g)$ - i.e., combine the earlier size and composition of government models

Policy

$$
\mathbf{q}=\left[\tau, g,\left\{f^{J}\right\}, r\right] \quad \Sigma_{J} f^{J}+g+r=3 \tau
$$

$r$ endogenous rent extracted by politicians

Rich, three-dimensional conflict over $\mathbf{q}$
among voters: over $\left\{f^{J}\right\}$
between voters and politicians: over $\tau, g$ vs. $r$ among politicians: over $r$

Basic hypothesis
resolution of these conflicting interests over $\mathbf{q}$ depends systematically on constitutional rules
cf. IO/micro theory, where resolution of conflicting interests between firms and consumers depend on market regulation rules

## B. Electoral rules and political competition

Electoral rules: two basic aspects
(i) electoral formula: vote shares $\rightarrow$ seat shares? ) winner takes all with plurality rule - proportionality with PR
(ii) district magnitude: \# of lawmakers elected in average district? typically low (often 1 ) with plurality rule - higher with PR
a third is ballot structure: voting for individuals or party lists? are lists are open or closed - can preference votes be cast?
(i) and (ii) quite strongly correlated in real-world systems, but not perfectly so

## 1. Adapt probabilistic voting model of Lecture 1

Traditional electoral competition two parties (candidates), $C=A, B$
commitment to platforms before election, here $\mathbf{q}_{C}$ ( $r_{C}$ implicit) opportunistic objectives, maximize

$$
\begin{equation*}
E\left(v_{C}\right)=p_{C}\left(R+\gamma r_{C}\right) \tag{3}
\end{equation*}
$$

$R$ exogenous ego-rent, $p_{C}$ probability $C$ wins the election, $(1-\gamma)$ transaction cost

Candidate preferences
$i$ in $J=1,2,3$ votes for $A$ if

$$
\begin{equation*}
W^{J}\left(\mathbf{q}_{A}\right)>W^{J}\left(\mathbf{q}_{B}\right)+\sigma^{i, J}+\delta \tag{4}
\end{equation*}
$$

Individual preferences: assume, not quite as before
$\sigma^{i, J} \lesseqgtr 0$ have group-specific distribution uniform on $\left[-\frac{1}{2 \phi^{J}}+\bar{\sigma}^{J}, \frac{1}{2 \phi^{J}}+\bar{\sigma}^{J}\right]$
$\phi^{2}>\phi^{1}, \phi^{3}, \quad \bar{\sigma}^{1}<\bar{\sigma}^{2}=0<\bar{\sigma}^{3}, \bar{\sigma}^{1} \phi^{1}+\bar{\sigma}^{3} \phi^{3}=0$
generate group biases for each party + convenient normalization
Aggregate popularity: assume (as before)

$$
\delta \lesseqgtr 0 \text { uniform on }\left[-\frac{1}{2 \psi}, \frac{1}{2 \psi}\right]
$$

Timing
parties know $\left\{\phi^{J}, \bar{\sigma}^{J}\right\}$ and $\psi$, but not $\delta$, when set $\mathbf{q}_{A}$
Swing voters
in group $J$ indifferent between $A$ and $B$

$$
\begin{equation*}
\sigma^{J}=W^{J}\left(\mathbf{q}_{A}\right)-W^{J}\left(\mathbf{q}_{B}\right)-\delta \tag{5}
\end{equation*}
$$

Vote share of $A$ in $J$
all voters with $\sigma^{i, J} \leq \sigma^{J}$ cast ballot for $A$

$$
\begin{equation*}
\pi_{A}^{J}=\phi^{J}\left(\sigma^{J}-\bar{\sigma}^{J}+\frac{1}{2 \phi^{J}}\right)=\frac{1}{2}+\phi^{J}\left(\sigma^{J}-\bar{\sigma}^{J}\right) \tag{6}
\end{equation*}
$$

depends on policy via (5)

How do parties evaluate policy?
$\mathbf{q}$ affects $E\left\{\pi_{A}^{J}\right\}$ via identity of swing voters in each $J$ parties trade off votes against votes (choice of $\tau, g,\left\{f^{J}\right\}$ ) and votes against rents (choice of $r$ )

But ...
how $p_{A}$ varies with $E\left\{\pi_{A}^{J}\right\}$ depends on electoral rule

## 2. PR elections (or large district magnitude)

Probability of winning
3 equal-sized districts $J$, same as groups: $f^{J}$ regional transfer enough seats and $\mathrm{PR} \Rightarrow A$ seat share is $s_{A}^{J}=\pi_{A}^{J}$ winning requires $50 \%$ of national vote $\pi_{A} \equiv \frac{1}{3} \Sigma \pi_{A}^{J} \geq \frac{1}{2}$

$$
\begin{equation*}
p_{A}=\operatorname{Prob}_{\delta}\left[\pi_{A} \geq \frac{1}{2}\right]=\frac{1}{2}+\frac{\psi}{3 \phi}\left[\Sigma \phi^{J}\left(W^{J}\left(\mathbf{q}_{A}\right)-W^{J}\left(\mathbf{q}_{B}\right)\right)\right] \tag{7}
\end{equation*}
$$

by (5)-(6), distributional assumptions, and $\phi \equiv \frac{1}{3} \Sigma \phi^{J}$
(alternative interpretation: plurality rule in 1 national district)
Equilibrium - establish a benchmark for comparison each party chooses identical policy platforms $\mathbf{q}_{A}=\mathbf{q}_{B}$ characterize equilibrium via four trade-offs associated with FOC for each policy instrument
(i) Redistributive transfers: $\left\{f^{J}\right\}$, votes in different groups? chase most responsive voters: $\phi^{2}>\phi^{1}, \phi^{3}$ complementary slackness: only $f^{2}>0$
hinges on quasi-linear utility (and uniform distribution)
(ii) Size of government: $f^{2}$ vs. $\tau$, votes in different groups? trade off group 2 votes against votes in all groups: because

$$
\phi^{2} \cdot 1>\sum_{J} \phi^{J} \cdot \frac{1}{3}=\phi
$$

$\tau=1$ optimal (with distortionary taxes $\tau<1$ ), as $\phi^{2}>\phi$
(iii) Public goods: $f^{2}$ vs. $g$, votes in different groups?
similar trade-off pins down supply of $g$

$$
\phi^{2} \cdot 1=\sum_{J} \phi^{J} \cdot H_{g}(g)
$$

note underprovision $H_{g}=\frac{\phi^{2}}{3 \phi}>\frac{1}{3}$ (utilitarian optimum) more pronounced if $\phi^{2}$ higher (more group-2 swing voters)
(iv) Rents: $r$, rents vs. votes (via lower $f^{2}$ )?
equilibrium rents may well be positive

$$
\frac{\gamma}{2}=p_{A} \gamma \leqslant-[R+\gamma r] \cdot \frac{d p_{A}}{d r}=[R+\gamma r] \cdot \frac{\psi \phi^{2}}{3 \phi} \quad[r \gg]
$$

( $p_{A}=\frac{1}{2}$ in equilibrium), higher if $\frac{\phi^{2}}{\phi}, \psi, R,(1-\gamma)$ lower

## 3. Plurality elections (or small district magnitude)

Probability of winning
again (now), 3 districts $J=1,2,3$, same as groups
by plurality, winner takes all, $s_{A}^{J}=1$ if $\pi_{A}^{J} \geq \frac{1}{2}$
if $-\bar{\sigma}^{1}$ and $\bar{\sigma}^{3}$ "large enough" $\Rightarrow$
$A$ and $B$ win their "safe districts", 1 and 3 , for sure
Equilibrium
$A, B$ only compete for $J=2$, where most swing voters

$$
p_{A}=\operatorname{Prob}_{\delta}\left[\pi_{A}^{2} \geq \frac{1}{2}\right]=\frac{1}{2}+\psi\left[W^{2}\left(\mathbf{q}_{A}\right)-W^{2}\left(\mathbf{q}_{B}\right)\right]
$$

(i) Redistributive transfers
$A, B$ incentives propose more $f^{2}$ than under PR
benefits the same, costs lower (ignored in district 1 and 3 )
(ii) Taxes
still optimal to set $\tau=1$
(iii) Public goods
further underprovision of $g$ than under PR , since

$$
\phi^{2} \cdot 1=\phi^{2} \cdot H_{g}(g)
$$

internalize benefits only for more narrow group of voters

$$
H_{g}(g)=1>\frac{\phi^{2}}{3 \phi}=\frac{\phi^{2}}{\Sigma \phi^{J}}
$$

(iv) Rents

$$
\frac{\gamma}{2}=p \gamma \leqslant-[R+\gamma r] \cdot \frac{d p_{A}}{d r}=[R+\gamma r] \cdot \psi
$$

competition now more fierce, i.e., $-\frac{d p_{A}}{d r}$ higher as $\psi>\psi \frac{\phi^{2}}{3 \phi}$ punishment for inefficiency larger than under PR

## 4. Comparative politics results

Plurality rule vs. PR (or small vs. large district magnitude)
focus competition to districts with many swing voters $\Rightarrow$

- more targeted redistribution (higher, concentrated $f$ )
- less public goods (lower $g$ )
- less rents (lower $r$ )
- same size of government (same $\tau$ )

Compare to predictions in other models

> P-T (2000, Ch. 9)
result on $g$ vs. $f$, and $\tau$ the same, result on $r$ opposite
L-P (2001) result on $g$ vs. $f$ the same
M-P-R (2002)
result on $g$ vs. $f$ the same, $\tau$ lower with majoritarian elections

## C. Forms of government and legislative bargaining

Parliamentary and presidential regimes
differ both in executive-legislature relations and rules for legislation

Confidence requirements of legislature
parliamentary: executive accountable to legislature presidential: no confidence requirement - as the executive (president) typically directly elected by voters

Separation of powers over legislation
parliamentary: powers often concentrated in cabinet presidential: powers often separated across offices and legislators

1. Combine retrospective voting and legislative bargaining

Use earlier building blocks
policy problem in A, legislative-bargaining model in Lecture 4, retrospective-voting/political-agency model in Lecture 4

Legislators
3 incumbents $J$ with opportunistic objectives

$$
E\left(v_{J}\right)=\gamma r_{J}+p_{J} R,
$$

Voters
in each district $J$, use retrospective voting rules

$$
p_{J}=\left\{\begin{array}{cc}
1 & \text { if } W^{J}(\mathbf{q}) \geq \varpi^{J} \\
& 0 \text { otherwise }
\end{array}\right.
$$

coordination within, but not across, districts

## 2. A simple legislature

Unrealistic warm up neither confidence requirement, nor separation of powers

Timing
see Figure 1
Equilibrium conditions
(i) $\mathbf{q}(\varpi)$ dominates $\overline{\mathbf{q}}$ for at least one legislator $J \neq A$, any $\varpi$
(ii) $\mathbf{q}(\varpi)$ optimal for $A$, any $\varpi$, given (i)
(iii) $\varpi^{J}$ optimal for voters in $J$, given (i), (ii) and $\varpi^{-J}$


Equilibrium policy

$$
\begin{align*}
r^{\mathrm{S}} & =r_{A}=3-\bar{r}-\frac{R}{\gamma} \\
\tau^{\mathrm{S}} & =1  \tag{8}\\
f^{\mathrm{S}} & =f^{A}=\bar{r}+\frac{R}{\gamma}-g \\
H_{g}\left(g^{\mathrm{S}}\right) & =1
\end{align*}
$$

Sketch of proof
consider arbitrary $\varpi=\left(\varpi^{A}, \varpi^{I}, \varpi^{K}\right)>y+1-\bar{r}$ and $A$ seeking re-election for herself and only one more $J$

Incentive constraint
joint "Leviathan" deviation should not pay

$$
\begin{equation*}
\gamma r+R \geq \gamma(3-\bar{r}) \tag{9}
\end{equation*}
$$

as re-election good enough for $X(R>\gamma \bar{r}) \Rightarrow r_{J}=0, J \neq A$

Optimal behavior for $A$
pick $I=X$ if $\varpi^{I}<\varpi^{K}$, assume this is the case
$\operatorname{Max} r_{A} \Rightarrow \tau=1, f^{K}=0$,

$$
W^{I}(\mathbf{q}(\varpi))=\varpi^{I}, W^{A}(\mathbf{q}(\varpi))=\varpi^{A}
$$

nail lawmaker $I$ and voters in $A, I, K$ to lowest possible payoff
Utility of voters
$W^{A}(\mathbf{q}(\varpi))=y+f^{A}+H(g)=\left(\varpi^{A}-H(g)\right)+H(g)$
$W^{I}(\mathbf{q}(\varpi))=y+f^{I}+H(g)=\left(\varpi^{I}-H(g)\right)+H(g)$
$W^{K}(\mathbf{q}(\varpi))=y+H(g)$
$W^{I}$ and $W^{K}$ discontinuous at $\varpi^{I}=\varpi^{K}$ as long as $f^{I}>0$
voters in $I$ and $K$ compete for any policy favors $\Rightarrow f^{I}=f^{K}=0$ only lawmaker $A$ has a binding re-election constraint

Voters in $A$
set $\varpi^{A}$ equal to solution of

$$
\begin{gathered}
\operatorname{Max}_{\tau, f} f^{A}, g \\
\\
\text { s t } \quad f^{A}+g \leq \bar{r}+\frac{R}{\gamma}+3(\tau-1)
\end{gathered}
$$

where constraint combines (2) and (9)

Optimal policy
voters in $A$ agree with legislator $A$ on $\tau=1$ want to set $H_{g}(1)=1$, trade off $g$ and $f^{A}$ one for one induce legislator $A$ to give up minimum rents

Three political failures (from the voters' horizon)
powerful $A$ gets large benefits for her district
waste via positive equilibrium rents
underprovision of public goods

## 3. Presidential-congressional regime

Crucial features
separation of powers, but no confidence requirement (cf. US)
Timing
Figure 2
$A_{\tau}, A_{g}$ control different instruments, sequential decisions


Equilibrium policy

$$
\begin{align*}
r^{\operatorname{Pr}} & =0 \\
\tau^{\operatorname{Pr}} & =\in\left[\frac{g}{3}, \frac{1}{3}\left(\bar{r}+\frac{R}{\gamma}\right)\right]  \tag{10}\\
f^{\mathrm{Pr}} & =f^{A_{g}} \in\left[0, \bar{r}+\frac{R}{\gamma}-g\right] \\
H_{g}\left(g^{\mathrm{Pr}}\right) & =1
\end{align*}
$$

only give intuition for these results
Spending stage
(i) competition among voters in $J \neq A_{g}$ drives transfers to 0
(ii) $A_{g}$ can please her voters, given available $\tau$,
$\Rightarrow$ trade off $f^{A_{g}}$ and $g$ one for one, so underprovision of $g$
(iii) as $\tau$ is given, voters in $A_{g}$ can insist on $r=r_{A}=0$

## Taxation stage

(i) $A_{\tau}$ and voters in $A_{\tau}$ not residual claimants on $\tau$
(ii) may have to give up some revenue to voters in $A_{g}$ (so have multiple equilibria)

Overall conclusion
separation of powers and lack of stable majority $\Rightarrow$
$f$ is targeted to minority, get small broad program $g$,
$r$ and $\tau$ are kept minimal

## 4. Parliamentary regime

Crucial features
concentration of powers to cabinet
confidence requirement: powers are maintained only if the cabinet survives (cf. UK)

Timing
Figure 3
Incentives for stable majority
veto right for government members $A_{\tau}, A_{g}$, costly to exercise
confidence vote followed by a government crisis this creates "legislative cohesion"
Figure 3


Table 2.1
Constitutions and economic policy
Theoretical predictions

Policy outcome

> Electoral rules
> Majoritarian vs.
> proportional

Form of government
Presidential vs. parliamentary

Overall size of government

Composition: broad vs. narrow programs
$-/ ?$
-


Equilibrium policy

$$
\begin{align*}
r^{\mathrm{Pa}} & =3-\frac{2 R}{\gamma} \\
\tau^{\mathrm{Pa}} & =1  \tag{11}\\
f^{\mathrm{Pa}} & =f^{A_{g}}+f^{A_{\tau}}>0 \\
\frac{1}{2} & \leq H_{g}\left(g^{\mathrm{Pa}}\right)<1
\end{align*}
$$

again only give intuition

## Spending

(i) cost of break-up: both groups of voters backing government have some bargaining power
(ii) benefits of $g$ internalized by majority, this gives less under-provision than in $\mathbf{2}$ and $\mathbf{3}$
(iii) outcome jointly optimal for $J=A_{\tau}, A_{g}$ but can split benefits in different ways (multiple equilibria)

## Rents

$r$ high, as politicians can collude, absent separation of powers

## Taxes

voters backing government want high $\tau$ as they, and their legislators, are residual claimants on revenue

## 5. Comparative politics results

Presidential vs. parliamentary regimes
no confidence requirement, but separation of powers $\Rightarrow$

- redistribution more targeted ( $f$ to one group)
- less public goods (lower $g$ )
- less rents (lower $r$ )
- smaller size of government (lower $\tau$ )

Compare to predictions in other work
C-K (2011) on US cities find analogous result for $\tau$

## D. Summary of predictions

Theoretically predicted effects
size of government spending
composition into broad vs. targeted programs (rents extracted by politicians)

+ or - in Table 2.1

How take these to data? Do the predictions hold up?
those are the topics in last lecture

Table 2.1
Constitutions and economic policy
Theoretical predictions

Policy outcome

> Electoral rules
> Majoritarian vs.
> proportional

Form of government
Presidential vs. parliamentary

Overall size of government

Composition: broad vs. narrow programs
$-/ ?$
-


## Empirical Strategy and Results Introduction

How do we test predictions in last lecture as summarized in Table 2.1

Describe data, methods, and selected results
Agenda
A. Data and their properties
B. Statistical concerns
C. Identification: effects on government size
D. Tests of other predictions

## A. Data and their properties

Sample selection
given questions: most interesting variation across countries limit sample to democracies (in retrospect, a mistake!) treat democracy status as random generous definition of democracy, test if matters

Two data sets
average annual obs. in 1990s for 85 democracies (cross section) annual obs. in 1960-98 for 60 democracies (panel)

Policy and performance measures
for each question in Table 2.1
describe in context

Country characteristics
many socioeconomic, historical, geographical, cultural variables covariates with policy or constitutional rules

Table 2.1
Constitutions and economic policy
Theoretical predictions

Policy outcome

> Electoral rules
> Majoritarian vs.
> proportional

Form of government
Presidential vs. parliamentary

Overall size of government

Composition: broad vs. narrow programs
$-/ ?$
-


Measures of electoral rules
binary indicator, MAJ, for electoral formula is only plurality rule used in election of lower house? yes: $M A J=1$, no: $M A J=0$

Measures of forms of government
binary indicator, $P R E S$, for confidence requirement is the executive independent of confidence in legislature? yes: $P R E S=1$, no: $P R E S=0$
no measure of separation of powers

Main characteristics
two features important for empirical strategy
(i) Constitutional inertia
deep reforms rare events: panel has no switch in $P R E S$
five in MAJ - ten if count mixed systems as well and several more in detailed measures of electoral rules must estimate constitutional effects from cross-country variation
(ii) Non-random selection of constitutional rules reflect history, geography and culture Figure 4.1 must be careful in inference

## Constitutional Atlas 1998



## B. Statistical concerns

Causal effects or statistical correlations?
serious pitfalls in inference from cross-sectional data try to address by applying alternative methods

What form of simultaneity
'reverse causation' perhaps not major issue (inertia over 40 years) 'omitted variables' more serious problem

How avoid confounding constitutions and other policy determinants? hold constant observables which are correlated with policy, outcomes as well as constitution selection $\Rightarrow$ use regression with many controls

How about remaining unobservables?
lacking imagination or difficulties in measurement can lead to bias isolate exogenous variation $\Rightarrow I V$-estimation clean estimates from selection bias $\Rightarrow$ Heckman-style adjustment

How about unwarranted extrapolation for heterogenous groups?
allow for non-linear relations in data and give higher weight to 'local' comparisons $\Rightarrow$
estimate with matching methods
Illustrate in the context of our problem use alternative methods to estimate same parameter and discuss specific identifying assumptions in theory and practice apply to constitutional effects on the size of government

## C. Identification: effects on government size

Details of empirical strategy for cross-sectional data what's the parameter of interest? under what assumptions can we estimate it, given concerns in $\mathbf{B}$ ? what results do we obtain?

Use size of government as specific example

## 1. Overall question and problem

Parameter of interest
what is direct effect of constitutional reform in country selected at random (ATE)?

How does rule $S=1$, vs. $S=0$, affect policy outcome $Y$ ?

$$
\begin{equation*}
\alpha \equiv \mathrm{E}\left(Y^{1}-Y^{0}\right)=\mathrm{E}\left\{E\left(Y^{1} \mid \mathbf{X}\right)\right\}-\mathrm{E}\left\{E\left(Y^{0} \mid \mathbf{X}\right)\right\} \tag{14}
\end{equation*}
$$

where last equality relies on law of iterated expectations when
$E$ and E refer to unobservables and observables, respectively
More concretely
how does switch $P R E S$ or $M A J$ from 0 to 1 affect central government spending (revenue) as \% of GDP?

Problem with observational data
need country- $i$ outcome in two potential states: $Y_{i}^{1}, Y_{i}^{0}$
observe only $Y_{i}=S_{i} Y_{i}^{1}+\left(1-S_{i}\right) Y_{i}^{0}$
other outcome is unobserved counterfactual

Pose problem differently
condition on $S$ rather than on $\mathbf{X}$, rewrite (14) as

$$
\begin{gather*}
\alpha=\mathrm{E}\left(Y^{1}-Y^{0}\right)=P \cdot\left[E\left(Y^{1} \mid S=1\right)-E\left(Y^{0} \mid S=1\right)\right]  \tag{15}\\
+(1-P) \cdot\left[E\left(Y^{1} \mid S=0\right)-E\left(Y^{0} \mid S=0\right)\right]
\end{gather*}
$$

$\alpha$ sums effects (ATN and $A T T)$ for countries in states $S=0,1$ which have probabilities $1-P$ and $P$, respectively
(Hypothetical) Experimental data
safe assume e.g., $\mathrm{E}\left(Y^{0} \mid S=1\right)=\mathrm{E}\left\{E\left(Y^{0} \mid \mathbf{X}, S=1\right)\right\}$
$=\mathrm{E}\left\{E\left(Y^{0} \mid \mathbf{X}, S=0\right)\right\}=\mathrm{E}\left(Y^{0} \mid S=0\right)$
because both unobserved and observed features balanced across $S=0,1$ - by randomization

Yet another way to pose the problem
consider general data-generating process

$$
\begin{gather*}
Y_{i}^{S}=F^{S}\left(\mathbf{X}_{i}\right)+\varepsilon_{i}^{S}, \quad S=0,1 \\
S_{i}=\left\{\begin{array}{c}
1 \text { as } G\left(\mathbf{W}_{i}, \mathbf{Z}_{i}\right)+\eta_{i} \geq 0 \\
0 \text { otherwise }
\end{array}\right. \tag{16}
\end{gather*}
$$

where $\mathbf{W}$ is subset of $\mathbf{X}$, while $\mathbf{Z}$ is not

Identification of $\alpha$
need strong explicit or implicit assumptions about $F^{S}(\mathbf{X}), \varepsilon^{S}, \eta$ to tackle the statistical problems mentioned in $\mathbf{B}$

## 2. Linear regression estimates

(i) Recursivity: $\operatorname{Cov}\left(\varepsilon^{S}, \eta\right)=0$
or, equivalently, conditional independence

$$
E\left(Y^{1} \mid \mathbf{X}, S=1\right)=E\left(Y^{1} \mid \mathbf{X}, S=0\right)=E\left(Y^{1} \mid \mathbf{X}\right)
$$

selection is random, after controlling for $\mathbf{X}$
$\mathbf{X}$ income, age and quality of democracy, openness, demographics, indicators for federal structure, OECD, continents, colonial history
(ii) Linearity

$$
Y_{i}^{S}=F^{S}\left(\mathbf{X}_{i}\right)+\varepsilon_{i}^{S}=\alpha^{S}+\beta \mathbf{X}_{i}+\varepsilon_{i}^{S}
$$

By (i) and (ii),

$$
\alpha=\alpha^{1}-\alpha^{0}=\mathrm{E}\left\{E\left(Y^{1}-Y^{0} \mid \mathbf{X}\right)\right\}
$$

causal effect estimated by coefficient on $S$ in OLS of $Y$ on $\mathbf{X}, S$

$$
Y_{i}=\alpha^{0}+\alpha S_{i}+\boldsymbol{\beta} \mathbf{X}_{i}+e_{i}
$$

where $e_{i}=\varepsilon_{i}^{0}+S_{i}\left(\varepsilon_{i}^{1}-\varepsilon_{i}^{0}\right)$

Results in Table 6.1

Table 6.1
Size of government and constitutions
Simple regression estimates

|  | (1) | (2) | (3) | (4) | (5) | (6) | (7) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Dep. var. | CGEXP | CGEXP | CGEXP | CGREV | CGEXP | CGEXP | CGEXP |
| PRES | $\begin{aligned} & -6.08 \\ & (1.97)^{* * *} \end{aligned}$ | $\begin{aligned} & -5.29 \\ & (1.92)^{* * *} \end{aligned}$ |  | $\begin{aligned} & -5.17 \\ & (2.44)^{* *} \end{aligned}$ | $\begin{aligned} & -8.29 \\ & (2.72)^{* * *} \end{aligned}$ | $\begin{aligned} & -3.46 \\ & (3.88) \end{aligned}$ | $\begin{aligned} & -7.49 \\ & (2.72)^{* * *} \end{aligned}$ |
| MAJ | $\begin{aligned} & -3.29 \\ & (1.73)^{*} \end{aligned}$ | $\begin{aligned} & -5.74 \\ & (1.95)^{* * *} \end{aligned}$ |  | $\begin{aligned} & -3.03 \\ & (1.85) \end{aligned}$ | $\begin{aligned} & -5.59 \\ & (2.68)^{* *} \end{aligned}$ | $\begin{aligned} & -2.93 \\ & (3.09) \end{aligned}$ | $\begin{gathered} -4.81 \\ (2.75)^{*} \end{gathered}$ |
| PROPRES |  |  | $\begin{aligned} & -7.08 \\ & (2.70)^{* *} \end{aligned}$ |  |  |  |  |
| MAJPAR |  |  | $\begin{aligned} & -7.30 \\ & (3.02)^{* *} \end{aligned}$ |  |  |  |  |
| MAJPRES |  |  | $\begin{aligned} & -10.36 \\ & (2.70)^{* * *} \end{aligned}$ |  |  |  |  |
| Continents | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Colonies | No | Yes | Yes | Yes | Yes | Yes | Yes |
| Sample | 90s, broad | 90s, broad | 90s, broad | 90s, broad | 90s, narrow | 60-90s, broad | $\begin{aligned} & 90 \mathrm{~s}, \mathrm{obs} \\ & \text { as(6) } \end{aligned}$ |
| Obs. | 80 | 80 | 80 | 76 | 62 | 60 | 60 |
| Adj. R2 | 0.58 | 0.63 | 0.63 | 0.58 | 0.60 | 0.54 | 0.63 |

Summary
PRES $=1 \Rightarrow$ spending more than $5 \%$ of GDP smaller
$M A J=1 \Rightarrow$ slightly smaller effect
reflect more rapid growth of government, in 1965-85
constitutional effects appear additive
Is identification convincing ?
have we included all the relevant variables in $\mathbf{X}$ to rule out 'history and culture' determining both $S$ and $Y$ ?
can we trust there are no interaction effects or other non-linearities?
address these in turn in sections $\mathbf{3}$ and 4

## 3. Relax conditional independence

Can we rule out selection on unobservables ?

$$
\text { i.e., is } \operatorname{Cov}(e, \eta) \neq 0 \text { ? }
$$

Two prospective sources of selection bias

$$
\begin{aligned}
& \quad p \lim (\hat{\alpha})=\alpha+\mathrm{E}\left(\varepsilon^{1} \mid S=1\right)-\mathrm{E}\left(\varepsilon^{0} \mid S=0\right)= \\
& \alpha+\left\{\mathrm{E}\left(\varepsilon^{0} \mid S=1\right)-\mathrm{E}\left(\varepsilon^{0} \mid S=0\right)\right\}+\mathrm{E}\left(\varepsilon^{1}-\varepsilon^{0} \mid S=1\right)
\end{aligned}
$$

conventional omitted variables: $\varepsilon^{0}$ and $S$ correlated heterogenous constitutional effects: $\varepsilon^{1}-\varepsilon^{0}$ and $S$ correlated

How relax conditional independence?
a. use instrumental variables to isolate exogenous variation in $S$
b. adjust estimates, Heckman-style, for correlation $e, \eta$

## a. Instrumental variables

Identifying assumptions
how find $\mathbf{Z}$ such that $\operatorname{Cov}(\mathbf{Z}, S) \neq 0$ and $\operatorname{Cov}(\mathbf{Z}, e)=0$ ?
timing of latest constitutional reform (3 indicator variables)
historical waves of reform (hold constant age of democracy)
latitude, fractions speaking English, European language
geographic and cultural 'distance' to old democratic institutions
Relevant?
constitutional timing: yes, weakly
distance measures: yes, definitely
Exogenous?
constitutional timing: yes, a priori distance measures: less certain
can test over-identifying restrictions, but low power

Apply to our data - results Table 6.2

Compare to earlier estimates in Table 6.1
point estimates agree pretty well note parsimonious 1st stage (weak instruments) standard errors grow as 2nd stage is richer can't reject over-identifying assumptions

Table 6.2
Size of government and constitutions Heckman and Instrumental Variables estimates

|  | $(3)$ | $(4)$ |
| :--- | :--- | :--- |
| Dep. var. |  |  |
| PRES | CGEXP | CGEXP |
|  |  |  |
| MAJ | -8.65 | -4.50 |
|  | $(3.63)^{* *}$ | $(3.89)$ |
| Conts \& Cols | -3.90 | -5.12 |
|  | $(3.46)$ | $(3.61)$ |
| Sample |  |  |
| Endogenous | No | COL_UKA, |
| selection | $90 s$, broad | LAAM |
| Estimation | $P R E S$ | $90 \mathrm{~s}, \mathrm{broad}$ |
|  | $M A J$ | $P R E S$ |
| rho | $2 S L S$ | $M A J$ |
| Chi-2: over-id |  | $2 S L S$ |
| Adj. R2 | 4.64 |  |
| Obs. | 0.59 | 3.61 |

## b. Heckman-style adjustment

Well-known idea
estimate selection equation (probit or logit) corresponding to

$$
S_{i}=\left\{\begin{array}{c}
1 \text { as } G\left(\mathbf{W}_{i}, \mathbf{Z}_{i}\right)+\eta_{i} \geq 0 \\
0 \text { otherwise }
\end{array}\right.
$$

correct estimates of $\alpha$ for remaining correlation $\rho=\operatorname{Corr}(\eta, e)$

Identifying assumptions
same exclusion restrictions on $\mathbf{Z}$, as in $\mathbf{a}$ to avoid relying (only) on functional form

Apply to our data - results Table 6.2

Compare to estimates in Tables 6.1, 6.2
point estimates, if anything, more negative estimated $\rho>0$ (for $P R E S$ ) $\Rightarrow$ OLS had positive bias could allow joint selection of $M A J$ and PRES, or separate distributions for $\varepsilon^{1}, \varepsilon^{0}$ (heterogenous treatment effect)

Table 6.2
Size of government and constitutions Heckman and Instrumental Variables estimates

|  | (1) | (2) | (3) | (4) |
| :---: | :---: | :---: | :---: | :---: |
| Dep. var. | CGEXP | CGEXP | CGEXP | CGEXP |
| PRES | $\begin{aligned} & -10.50 \\ & (3.98)^{* * *} \end{aligned}$ | $\begin{aligned} & -5.37 \\ & (2.19)^{* *} \end{aligned}$ | $\begin{aligned} & -8.65 \\ & (3.63)^{* *} \end{aligned}$ | $\begin{aligned} & -4.50 \\ & (3.89) \end{aligned}$ |
| MAJ | $\begin{aligned} & -5.69 \\ & (1.86)^{* * *} \end{aligned}$ | $\begin{aligned} & -4.92 \\ & (2.57)^{*} \end{aligned}$ | $\begin{aligned} & -3.90 \\ & (3.46) \end{aligned}$ | $\begin{aligned} & -5.12 \\ & (3.61) \end{aligned}$ |
| Conts \& Cols | Yes | Yes | No | $\begin{aligned} & \text { COL_UKA, } \\ & \text { LAAM } \end{aligned}$ |
| Sample | 90s, broad | 90s, broad | 90s, broad | 90s, broad |
| Endogenous selection | PRES | MAJ | PRES <br> MAJ | PRES <br> MAJ |
| Estimation | Heckman 2-step | Heckman 2-step | 2SLS | 2SLS |
| rho | 0.64 | -0.02 |  |  |
| Chi-2: over-id Adj. R2 |  |  | 4.64 0.59 | 3.61 0.60 |
| Obs. | 75 | 75 | 75 | 75 |

## 4. Relax linearity

Many reasons believe $Y^{S}=F^{S}(\mathbf{X})$ non-linear
no great concern if $\mathbf{X}^{1}$ and $\mathbf{X}^{0}$ have similar distribution if not, specification bias - selection on observables - can be severe

Are $\mathbf{X}^{1}$ and $\mathbf{X}^{0}$ similar ?
test $E\left(\mathbf{X}^{1}\right)=E\left(\mathbf{X}^{0}\right)$ suggests not, cf. Table 5.3
PRES: reject in 7 cases out of 9
$M A J$ : reject in 4 cases out of 9
How take care of prospective specification bias ?
parsimonious assumption on functional form rely on 'local' comparisons

Table 5.3

## Balancing property

Equal-means tests for different constitutional groups

|  | Whole <br> sample |
| :--- | :--- |
|  |  |
| MAJ $=1$ vs. |  |
| MAJ $=0$ |  |
| LYP | 0.04 |
| PROP65 | 0.01 |
| GASTIL | 0.08 |
| FEDERAL | 0.93 |
| COL_UKA | 0.00 |
| LAAM | 0.34 |
| TRADE | 0.44 |
| PROT80 | 0.94 |
| CATHO80 | 0.00 |
|  |  |
| PRES $=1$ vs. |  |
| PRES $=0$ | 0.00 |
| LYP | 0.00 |
| PROP65 | 0.00 |
| GASTIL | 0.22 |
| FEDERAL | 0.44 |
| COL_UKA | 0.00 |
| LAAM | 0.01 |
| TRADE | 0.03 |
| PROT80 | 0.00 |
| CATHO80 |  |

Probabilities of falsely rejecting the hypothesis of equal means across constitutional groups under the hypothesis of equal variances.

Central idea in matching: mimic experimental measurement split data in 'treated' and 'controls' counterfactual for treated: controls with similar $\mathbf{X}_{i}$ estimate treatment effect of $S$ on $Y$ non-parametrically

Difficulty
too data-hungry, dimension of $\mathbf{X}$ large
Resolution
match on propensity score, rather than directly on $\mathbf{X}$
$p_{i}=p\left(\mathbf{X}_{i}\right)=\operatorname{Prob}\left[S=1 \mid \mathbf{X}_{i}\right]$

Identification
(i) conditional independence of $Y$ given $\mathbf{X}$
(ii) common-support condition: $0<p\left(\mathbf{X}_{i}\right)<1$, all $\mathbf{X}_{i}$ can rewrite (15) as

$$
\begin{gather*}
\alpha=P \cdot \mathrm{E}\left\{\left[E\left(Y^{1} \mid p\right)-E\left(Y^{0} \mid p\right)\right] \mid S=1\right\}+ \\
(1-P) \cdot \mathrm{E}\left\{\left[E\left(Y^{1} \mid 1-p\right)-E\left(Y^{0} \mid 1-p\right)\right] \mid S=0\right\} \tag{17}
\end{gather*}
$$

i.e., condition on $p$ and thus indirectly, not directly, on $X$

Number of practical questions in evaluating $\alpha$

1. How estimate $p\left(\mathbf{X}_{i}\right)$ ?
simple logit (or probit)
specification of $\mathbf{X}$ reflect concern for (trade-off between) conditional-independence and common-support conditions highest $t$-statistics in equal-means tests and regressions
2. Does matching indeed balance the observations?

Equal-means, for same $\mathbf{X}$ as before, in three strata for $p$ PRES: now reject only in 2 out of 27 cases (cf. Table 5.3) $M A J$ : again, reject in 2 out 27 cases
3. How estimate $P$ ?
relative sample frequency of $S=1$

Table 5.3

## Balancing property

Equal-means tests for different constitutional groups

|  | Whole <br> sample | $\mathrm{p}<0.33$ | $0.33<\mathrm{p}<0.67$ | $0.67<\mathrm{p}$ |
| :--- | :--- | :--- | :--- | :--- |
|  |  |  |  |  |
| MAJ=1 vs. |  |  |  |  |
| MAJ=0 |  |  |  |  |
| LYP | 0.04 | 0.04 | 0.62 | 0.21 |
| PROP65 | 0.01 | 0.32 | 0.90 | 0.04 |
| GASTIL | 0.08 | 0.33 | 0.55 | 0.37 |
| FEDERAL | 0.93 | 0.79 | 0.57 | 0.48 |
| COL_UKA | 0.00 | 0.69 | 0.42 | 0.35 |
| LAAM | 0.34 | 0.27 | 0.39 | 0.17 |
| TRADE | 0.44 | 0.13 | 0.93 | 0.31 |
| PROT80 | 0.94 | 0.56 | 0.75 | 0.37 |
| CATHO80 | 0.00 | 0.11 | 0.46 | 0.83 |
|  |  |  |  |  |
| PRES=1 vs. |  |  |  |  |
| PRES=0 |  | 0.87 | 0.01 | 0.54 |
| LYP | 0.00 | 0.34 | 0.39 | 0.86 |
| PROP65 | 0.00 | 0.59 | 0.22 | 0.71 |
| GASTIL | 0.00 | 0.07 | 0.30 | 0.27 |
| FEDERAL | 0.22 | 0.88 | 0.56 | 0.83 |
| COL_UKA | 0.44 | 0.53 | 0.23 | 0.22 |
| LAAM | 0.00 | 0.33 | 0.34 | 0.40 |
| TRADE | 0.01 | 0.65 | 0.60 | 0.22 |
| PROT80 | 0.03 | 0.28 | 0.24 | 0.02 |
| CATHO80 | 0.00 |  |  |  |

Probabilities of falsely rejecting the hypothesis of equal means across constitutional groups under the hypothesis of equal variances.
Strata defined on the common support of propensity scores, p, estimated by logit regressions including: LYP, PROP65, GASTIL, FEDERAL, COL_UKA, LAAM.
4. How estimate $\mathrm{E}\left\{E\left(Y^{1} \mid p\right) \mid S=1\right\}$ ?
sample mean among "treated"
5. How estimate $\mathrm{E}\left\{E\left(Y^{0} \mid p\right) \mid S=1\right\}$ ?
which controls matched with given $p_{i}$ among treated ?
(i) nearest-neighbor: the (one) control with closest $p_{i}$ produces (mostly) natural matches, cf. Table 5.2
(ii) stratification: arithmetic mean, all controls in same $p$ interval (iii) kernel: geometric mean, all controls in radius of $p_{i}$
6. How impose common-support condition (comparability) ? compute $3-5$ only for overlapping support of $p_{i}$

Table 5.2
Estimated propensity scores
(a) Majoritarian elections

| Country | PSCORE | MAJ | Country | PSCORE | MAJ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Uruguay | 0.052 | 0 | Nepal | 0.337 | 1 |
| Sweden | 0.070 | 0 | South Korea | 0.355 | 0 |
| Greece | 0.073 | 0 | Bangladesh | 0.371 | 1 |
| Bulgaria | 0.075 | 0 | Philippines | 0.377 | 1 |
| Italy | 0.077 | 0 | Namibia | 0.419 | 0 |
| UK | 0.078 | 1 | Barbados | 0.496 | 1 |
| Romania | 0.083 | 0 | New Zeland | 0.568 | 1 |
| Peru | 0.084 | 0 | Jamaica | 0.582 | 1 |
| Belgium | 0.090 | 0 | Ireland | 0.617 | 0 |
| Norway | 0.090 | 0 | Canada | 0.641 | 1 |
| France | 0.093 | 1 | Singapore | 0.659 | 1 |
| Spain | 0.095 | 0 | Israel | 0.673 | 0 |
| Latvia | 0.101 | 0 | Sri Lanka | 0.674 | 0 |
| Portugal | 0.104 | 0 | Trinidad\&Tobago | 0.694 | 1 |
| Denmark | 0.105 | 0 | Australia | 0.735 | 1 |
| Hungary | 0.106 | 0 | South Africa | 0.757 | 0 |
| Japan | 0.108 | 1 | Cyprus (G) | 0.759 | 0 |
| Colombia | 0.112 | 0 | Malta | 0.760 | 0 |
| Estonia | 0.114 | 0 | Bahamas | 0.763 | 1 |
| Guatemala | 0.115 | 0 | Pakistan | 0.781 | 1 |
| Czech Republic | 0.126 | 0 | Uganda | 0.790 | 1 |
| Luxembourg | 0.127 | 0 | Gambia | 0.794 | 1 |
| Chile | 0.128 | 1 | Ghana | 0.797 | 1 |
| Argentina | 0.132 | 0 | Zimbabwe | 0.808 | 1 |
| Finland | 0.132 | 0 | Belize | 0.812 | 1 |
| Paraguay | 0.133 | 0 | Fiji | 0.828 | 0 |
| Slovak Republic | 0.141 | 0 | Malawi | 0.831 | 1 |
| Nicaragua | 0.148 | 0 | St. Vincent\&Granada | 0.856 | 1 |
| Dominican Republic | 0.152 | 0 | Zambia | 0.856 | 1 |
| Netherlands | 0.153 | 0 | Malaysia | 0.857 | 1 |
| Ecuador | 0.157 | 0 | Mauritius | 0.873 | 1 |
| Germany | 0.160 | 0 | India | 0.886 | 1 |
| Russia | 0.161 | 0 | Papua New Guina | 0.904 | 1 |
| Poland | 0.177 | 0 | Botswana | 0.924 | 1 |
| Bolivia | 0.181 | 0 |  |  |  |
| Honduras | 0.185 | 0 |  |  |  |
| Mexico | 0.194 | 0 |  |  |  |
| Austria | 0.199 | 0 |  |  |  |
| Iceland | 0.212 | 0 |  |  |  |
| Switzerland | 0.214 | 0 |  |  |  |
| Turkey | 0.220 | 0 |  |  |  |
| Brazil | 0.230 | 0 |  |  |  |
| Costa Rica | 0.240 | 0 |  |  |  |
| El Salvador | 0.258 | 0 |  |  |  |
| Thailand | 0.264 | 1 |  |  |  |
| Venezuela | 0.292 | 0 |  |  |  |
| USA | 0.297 | 1 |  |  |  |
| Senegal | 0.320 | 0 |  |  |  |

PSCORE is the predicted value of a logit regression of MAJ on LYP, PROP65, FEDERAL, GASTIL, LAAM, COL_UKA
Boldface observations are discarded to impose common support.

Apply to our data - results Table 6.3
Compare to earlier estimates in Tables 6.1-6.2
point estimates agree
based on fewer observations on common support higher standard errors: trade off less bias against less efficiency

Summary: estimated constitutional effects consistent with theory
larger government, by $\sim 5 \%$ of GDP, of parliamentary democracy by about same amount of proportional democracy

Table 6.3
Size of government and constitutions
Matching estimates

|  | (1) | (2) | (3) | (4) | (5) | (6) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Dep. var. | CGEXP | CGEXP | CGEXP | CGEXP | CGEXP | CGEXP |
| PRES | $\begin{aligned} & -7.30 \\ & (2.30)^{* * *} \end{aligned}$ | $\begin{aligned} & -7.91 \\ & (2.90)^{* * *} \end{aligned}$ | $\begin{aligned} & -5.87 \\ & (4.93) \end{aligned}$ | $\begin{aligned} & -7.92 \\ & (5.11) \end{aligned}$ | $\begin{aligned} & -2.54 \\ & (2.30) \end{aligned}$ | $\begin{aligned} & -4.00 \\ & (3.45) \end{aligned}$ |
| MAJ | $\begin{aligned} & -5.76 \\ & (2.94)^{*} \end{aligned}$ | $\begin{aligned} & -6.55 \\ & (2.82)^{* *} \end{aligned}$ | $\begin{aligned} & -4.87 \\ & (3.65) \end{aligned}$ | $\begin{aligned} & -4.08 \\ & (4.16) \end{aligned}$ | $\begin{aligned} & -6.59 \\ & (3.06)^{* *} \end{aligned}$ | $\begin{aligned} & -8.81 \\ & (3.15)^{* * *} \end{aligned}$ |
| Estimation <br> Sample <br> Logit Specif. | Kernel 90s, broad 1 | Kernel 90s, broad 2 | Strat <br> 90s, broad <br> 1 | Strat <br> 90s, broad <br> 2 | Nearest 90s, broad 1 | Nearest 90s, broad 2 |
| Obs. on common support | $\begin{aligned} & 65 \text { PRES } \\ & 67 \mathrm{MAJ} \end{aligned}$ | $\begin{aligned} & 40 \text { PRES } \\ & 57 \text { MAJ } \end{aligned}$ | $\begin{aligned} & 65 \text { PRES } \\ & 67 \mathrm{MAJ} \end{aligned}$ | $\begin{aligned} & 40 \mathrm{PRES} \\ & 57 \mathrm{MAJ} \end{aligned}$ | $\begin{aligned} & 65 \text { PRES } \\ & 67 \mathrm{MAJ} \end{aligned}$ | $\begin{aligned} & 40 \mathrm{PRES} \\ & 57 \mathrm{MAJ} \end{aligned}$ |

## 5. Summary: size of government

Estimated constitutional effects
consistent with theory
larger government, by $\sim 5 \%$ of GDP, of parliamentary democracy by about same amount of proportional democracy

## D. Tests of other predictions

Quick description of results

## 1. Composition of government

Measurement
$Y$ welfare-state spending (social transfers), as \% of GDP
X same as for overall spending
Estimation
same battery of methods as in $\mathbf{C}$
Summary of findings
partly consistent with theory
$2-3 \%$ of GDP higher in (good, old) proportional democracies
$2-3 \%$ of GDP higher in (good, old) parliamentary democracies

## 2. Political rents and corruption

Measurement
$Y$ perception indexes for corruption, and government (in)effectiveness, 0-10 scale
X dozen covariates with corruption, suggested by earlier studies
Here can also exploit (piecemeal) electoral reforms
district magnitude \& ballot structure, as well as electoral formula cross-section and panel (fixed-effect) estimates agree

Summary of findings: electoral rules
larger electoral districts: less corruption (and inefficiency) more list voting: more corruption (and inefficiency) quantitatively important effects
binary indicator: no robust effect
Summary of findings: form of government
less corruption in presidential regimes?
perhaps, but only in better democracies

## 3. Extensions - examples of more recent research

(i) Deeper theoretical and empirical analysis mechanism for higher government spending under PR? incentives for politicians vs. indirect effect via party system
(ii) Wider scope of empirical work systematic effects beyond fiscal policy (and corruption)? trade policy, regulatory policy, economic performance
(iii) More extensive data sets with sharper identification from time variation around reforms? exploit switches in and out of democracy

