Political Economics II Spring 2018

Part III – Political Institutions and Economic Policy

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Introduction

Effects of political institutions – "constitutional reform"?
what happens to economic policy?
traditionally, neglected between economics and political science

Research in last fifteen years

new theory developed (applying tools you have seen already) and used as stepping stone for empirical analysis

Goal of Lectures 7-8

give selective introduction to research results and methods underlying theoretical models and predictions (Lecture 7) empirical measurement, strategy and results (Lecture 8)

Reforms of which constitutional provisions?

two reasonably fundamental provisions, where some theory exists electoral rules: "majoritarian" vs. "proportional" (and finer detail) forms of government: "presidential" vs. "parliamentary"

Which policies?

again, follow theory: fiscal policy (rent extraction) large variation in observed outcomes, even among similar units

Domain of application?

rules and policy outcomes at national level parallel literature about sub-national institutions and policy today, we begin with some theory and its predictions next week, we confront these predictions with data

Agenda of Lecture 7

- A. A public-finance problem
- B. Electoral rules and political competition
- C. Forms of government and legislative bargaining
- D. Summary of predictions

A. A public-finance problem

Groups of voters

homogenous (continuum) within J = 1, 2, 3, equal size, $N^J = 1$

Private preferences over economic outcomes

$$w^{J} = c^{J} + H(g) = y + (1 - \tau) + f^{J} + H(g)$$
 (1)

policies both targeted (f^J) and non-targeted (τ,g) – we combine the earlier size and composition of government models

Policy

$$\mathbf{q} = [\tau, g, \{f^J\}, r]$$

$$\Sigma_J f^J + g + r = 3\tau \tag{2}$$

r endogenous rent extracted by politicians

Rich, three-dimensional conflict over ${\bf q}$ among voters: over $\{f^J\}$ between voters and politicians: over τ,g vs. r among politicians: over r

Basic hypothesis

resolution of these conflicting interests over **q** depends systematically on constitutional rules

cf. IO/micro theory, where resolution of conflicting interests between firms and consumers depend on market regulation rules

B. Electoral rules and political competition

Electoral rules: two basic aspects

- (i) electoral formula: vote shares → seat shares? winner takes all with plurality rule – proportionality with PR
- (ii) district magnitude: # of lawmakers elected in average district? typically low (often 1) with plurality rule higher with PR
- a third is ballot structure: voting for individuals or party lists? are lists are open or closed can preference votes be cast?
- (i) and (ii) quite strongly correlated in real-world systems, but not perfectly so

1. Adapt probabilistic voting model of Lecture 1

Traditional electoral competition

two parties (candidates), C = A, Bcommitment to platforms before election, here \mathbf{q}_C (r_C implicit) opportunistic objectives, maximize

$$E(v_C) = p_C(R + \gamma r_C) \tag{3}$$

R exogenous ego-rent, p_C probability C wins the election, $(1 - \gamma)$ transaction cost

Candidate preferences among voters

i in J = 1, 2, 3 votes for A if

$$W^{J}(\mathbf{q}_{A}) > W^{J}(\mathbf{q}_{B}) + \sigma^{i,J} + \delta \tag{4}$$

Individual preferences: assume, not quite as before

$$\sigma^{i,J} \leq 0$$
 have group-specific distribution uniform on $\left[-\frac{1}{2\phi^J} + \overline{\sigma}^J, \frac{1}{2\phi^J} + \overline{\sigma}^J \right]$

(i)
$$\phi^2 > \phi^1, \phi^3$$
, (ii) $\overline{\sigma}^1 < \overline{\sigma}^2 = 0 < \overline{\sigma}^3$, (iii) $\overline{\sigma}^1 \phi^1 + \overline{\sigma}^3 \phi^3 = 0$ (i)-(ii) generate group biases for parties, (iii) for convenience

Aggregate popularity: precisely as in Lecture 3

$$\delta \leq 0$$
 uniform on $\left[-\frac{1}{2\psi}, \frac{1}{2\psi}\right]$

Timing

parties know $\{\phi^J, \overline{\sigma}^J\}$ and ψ , but not δ , when set \mathbf{q}_A

Swing voters

in group J indifferent between A and B

$$\sigma^{J} = W^{J}(\mathbf{q}_{A}) - W^{J}(\mathbf{q}_{B}) - \delta \tag{5}$$

Vote share of A in J

all voters with $\sigma^{i,J} \leq \sigma^J$ cast ballot for A

$$\pi_A^J = \phi^J(\sigma^J - \overline{\sigma}^J + \frac{1}{2\phi^J}) = \frac{1}{2} + \phi^J(\sigma^J - \overline{\sigma}^J) \tag{6}$$

depends on policy via (5)

How do parties evaluate policy?

q affects $E\{\pi_A^J\}$ via identity of swing voters in each J parties trade off votes against votes (choice of $\tau, g, \{f^J\}$) and votes against rents (choice of r)

But ...

how p_A varies with $E\{\pi_A^J\}$ depends on electoral rule

2. PR elections (or large district magnitude)

Probability of winning

3 equal-sized districts J, same as groups: f^J regional transfer enough seats and PR \Rightarrow A seat share is $s_A^J \simeq \pi_A^J$ winning requires 50% of national vote $\pi_A \equiv \frac{1}{3} \Sigma \pi_A^J \geq \frac{1}{2}$

$$p_A = \operatorname{Prob}_{\delta}[\pi_A \ge \frac{1}{2}] = \frac{1}{2} + \frac{\psi}{3\phi} \left[\sum \phi^J(W^J(\mathbf{q}_A) - W^J(\mathbf{q}_B)) \right] \quad (7)$$

by (5)-(6), distributional assumptions, and $\phi \equiv \frac{1}{3}\Sigma\phi^J$ (alternative interpretation: plurality rule in 1 national district)

Equilibrium – establish a benchmark for comparison

each party chooses identical policy platforms $\mathbf{q}_A = \mathbf{q}_B$ characterize equilibrium via four trade-offs – consider the FOC for each policy instrument

- (i) Redistributive transfers: $\{f^J\}$, expected votes in different groups? chase most responsive voters: $\phi^2 > \phi^1, \phi^3$ complementary slackness: only $f^2 > 0$ hinges on quasi-linear utility (and uniform distribution)
- (ii) Size of government: f^2 vs. τ , expected votes in different groups? trade off group 2 votes against votes in all groups: because

$$\phi^2 \cdot 1 > \sum_{J} \phi^J \cdot \frac{1}{3} = \phi$$

 $\tau = 1$ optimal (with distortionary taxes $\tau < 1$), as $\phi^2 > \phi$

(iii) Public goods: f^2 vs. g, expected votes in different groups? similar trade-off pins down supply of g

$$\phi^2 \cdot 1 = \sum_{J} \phi^J \cdot H_g(g)$$

note underprovision $H_g = \frac{\phi^2}{3\phi} > \frac{1}{3}$ (utilitarian optimum) more pronounced if ϕ^2 higher (more group-2 swing voters)

(iv) Rents: r vs f^2 , rents vs. expected votes (via lower f^2)? equilibrium rents may well be positive

$$\frac{\gamma}{2} = p_A \gamma \leqslant -[R + \gamma r] \cdot \frac{dp_A}{dr} = [R + \gamma r] \cdot \frac{\psi \phi^2}{3\phi} \qquad [r \geqslant 0]$$

 $(p_A = \frac{1}{2} \text{ in equilibrium}), \text{ higher if } \frac{\phi^2}{\phi}, \psi, R, (1 - \gamma) \text{ lower}$

3. Plurality elections (or small district magnitude)

Probability of winning

again (now), 3 districts J = 1, 2, 3, same as groups by plurality, winner takes all, $s_A^J = 1$ if $\pi_A^J \geq \frac{1}{2}$ if $-\overline{\sigma}^1$ and $\overline{\sigma}^3$ "large enough" \Rightarrow

A and B win their "safe districts", 1 and 3, for sure

Equilibrium

A, B only compete for J=2, where most swing voters

$$p_A = \text{Prob}_{\delta}[\pi_A^2 \ge \frac{1}{2}] = \frac{1}{2} + \psi[W^2(\mathbf{q}_A) - W^2(\mathbf{q}_B)]$$

(i) Redistributive transfers

A,B have incentives to propose more f^2 than under PR benefits the same, costs lower (ignored in district 1 and 3) (ii) Taxes still optimal to set $\tau = 1$

(iii) Public goods

further underprovision of g than under PR, since

$$\phi^2 \cdot 1 = \phi^2 \cdot H_g(g)$$

internalize benefits only for more narrow group of voters

$$H_g(g) = 1 > \frac{\phi^2}{3\phi} = \frac{\phi^2}{\Sigma \phi^J}$$

(iv) Rents

$$\frac{\gamma}{2} = p\gamma \leqslant -[R + \gamma r] \cdot \frac{dp_A}{dr} = [R + \gamma r] \cdot \psi$$

competition now more fierce, i.e., $-\frac{dp_A}{dr}$ higher as $\psi > \psi \frac{\phi^2}{3\phi}$ punishment for inefficiency larger than under PR

4. Comparative politics results

Plurality rule vs. PR (or small vs. large district magnitude) focus competition to districts with many swing voters \Rightarrow

- more targeted redistribution (higher, concentrated f)
- less public goods (lower g)
- less rents (lower r)
- same size of government (same τ)

Compare to predictions in other models

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Persson-Tabellini (2000, Ch. 9)
result on g vs. f, and \tau the same, result on r opposite
Lizzeri-Persico (2001) result on g vs. f the same
MilesiFeretti-Perotti-Rostagno (2002)
result on g vs. f the same, \tau lower with majoritarian elections
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C. Forms of government and legislative bargaining

Parliamentary and presidential regimes

differ both in executive-legislature relations and rules for legislation

Confidence requirements of legislature

parliamentary: executive accountable to legislature presidential: no confidence requirement – as the executive (president) typically directly elected by voters

Separation of powers over legislation

parliamentary: powers often concentrated in cabinet

presidential: powers often separated across offices and legislators

1. Combine retrospective voting and legislative bargaining

Use earlier building blocks

policy problem in **A**, legislative-bargaining model in Lecture **4**, retrospective-voting/political-agency model in Lecture **4**

Legislators

3 incumbents J with opportunistic objectives

$$E(v_J) = \gamma r_J + p_J R,$$

Voters

in each district J, use retrospective voting rules

$$p_J = \begin{cases} 1 & \text{if } W^J(\mathbf{q}) \ge \varpi^J \\ 0 & \text{otherwise} \end{cases}$$

coordination within, but not across, districts

2. A simple legislature

Unrealistic warm up

neither confidence requirement, nor separation of powers

Timing

see Figure 1

Equilibrium conditions

- (i) $\mathbf{q}(\boldsymbol{\varpi})$ dominates $\overline{\mathbf{q}}$ for at least one legislator $J \neq A$, any $\boldsymbol{\varpi}$
- (ii) $\mathbf{q}(\boldsymbol{\varpi})$ optimal for A, any $\boldsymbol{\varpi}$, given (i)
- (iii) ϖ^J optimal for voters in J, given (i), (ii) and ϖ^{-J}

Figure 1

Wether & A

Lesislat

7,5 (2) (3)

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Equilibrium policy

$$r^{\mathrm{S}} = r_{A} = 3 - \overline{r} - \frac{R}{\gamma}$$

$$\tau^{\mathrm{S}} = 1$$

$$f^{\mathrm{S}} = f^{A} = \overline{r} + \frac{R}{\gamma} - g$$

$$H_{g}(g^{\mathrm{S}}) = 1$$

Sketch of proof

consider arbitrary $\boldsymbol{\varpi} = (\varpi^A, \varpi^I, \varpi^K) > y + 1 - \overline{r}$ and A seeking re-election for herself and only one more J

Incentive constraint

joint "Leviathan" deviation should not pay

$$\gamma r + R \ge \gamma (3 - \overline{r}) \tag{9}$$

as re-election good enough for X $(R > \gamma \overline{r}) \Rightarrow r_J = 0, J \neq A$

Optimal behavior for A

pick I = X if $\varpi^I < \varpi^K$, assume this is the case

$$\operatorname{Max} r_A \Rightarrow \tau = 1, f^K = 0, W^I(\mathbf{q}(\boldsymbol{\varpi})) = \boldsymbol{\varpi}^I, W^A(\mathbf{q}(\boldsymbol{\varpi})) = \boldsymbol{\varpi}^A$$

nail lawmaker I and voters in A, I, K to lowest possible payoff

Utility of voters

$$W^{A}(\mathbf{q}(\boldsymbol{\varpi})) = y + f^{A} + H(g) = (\boldsymbol{\varpi}^{A} - H(g)) + H(g)$$

$$W^{I}(\mathbf{q}(\boldsymbol{\varpi})) = y + f^{I} + H(g) = (\boldsymbol{\varpi}^{I} - H(g)) + H(g)$$

$$W^{K}(\mathbf{q}(\boldsymbol{\varpi})) = y + H(g)$$

 W^I and W^K discontinuous at $\varpi^I = \varpi^K$ as long as $f^I > 0$ voters in I and K compete for any policy favors $\Rightarrow f^I = f^K = 0$ only lawmaker A has a binding re-election constraint

Voters in A

set ϖ^A equal to solution of

$$\operatorname{Max}_{\tau,f^A,g}[y+(1-\tau)+f^A+H(g)] ,$$
 s t $f^A+g \leq \overline{\tau}+\frac{R}{\gamma}+3(\tau-1)$

where constraint combines (2) and (9)

Optimal policy

voters of A agree with lawmaker A on $\tau = 1$ want to set $H_g(1) = 1$, trade off g and f^A one for one induce lawmaker A to give up minimum rents Three political failures (from the voters' horizon)

powerful lawmaker A gets large benefits for her district

waste via positive equilibrium rents

underprovision of public goods

3. Presidential-congressional regime

Crucial features

separation of powers, but no confidence requirement (cf. US)

Timing

Figure 2

 A_{τ}, A_{g} control different instruments, sequential decisions

SSANGAOS 4

Equilibrium policy

$$r^{\text{Pr}} = 0$$

$$\tau^{\text{Pr}} = \left[\frac{g}{3}, \frac{1}{3}(\overline{r} + \frac{R}{\gamma})\right]$$

$$f^{\text{Pr}} = f^{A_g} \in \left[0, \overline{r} + \frac{R}{\gamma} - g\right]$$

$$H_g(g^{\text{Pr}}) = 1$$

only give intuition for this result

Spending stage

- (i) competition among voters in $J \neq A_g$ drives transfers to 0
- (ii) lawmaker A_g can please her voters, given available τ ,
 - \Rightarrow trade off f^{A_g} and g one for one, so underprovision of g
- (iii) as τ is given, voters in A_g can insist on $r = r_A = 0$

Taxation stage

- (i) lawmaker A_{τ} and voters in A_{τ} not residual claimants on τ
- (ii) may have to give up some revenue to voters in A_g (so have multiple equilibria)

Overall conclusion

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separation of powers and lack of stable majority \Rightarrow f targeted to minority, small broad program g, r and \tau kept minimal
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4. Parliamentary regime

Crucial features

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concentration of powers to cabinet
confidence requirement: powers are maintained only
as long as cabinet survives (cf. UK)
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Timing

Figure 3

lawmakers A_{τ} , A_{g} control different instruments, bargain bilaterally

Incentives for stable majority

veto right for government members A_{τ}, A_{g} , is costly to exercise confidence vote followed by a government crisis this creates "legislative cohesion"

Election

n

1 zo zo!

Equilibrium policy

$$r^{\text{Pa}} = 3 - \frac{2R}{\gamma}$$

$$\tau^{\text{Pa}} = 1$$

$$f^{\text{Pa}} = f^{A_g} + f^{A_{\tau}} > 0$$

$$\frac{1}{2} \le H_g(g^{\text{Pa}}) < 1$$

$$(1)$$

again give only intuition

Spending

- (i) cost of break-up: both groups of voters back government have some bargaining power against each other
- (ii) benefits of g internalized by a majority, not a minority and this gives less under-provision than in 2 and 3
- (iii) outcome jointly optimal for groups $J = A_{\tau}, A_g$ but can split benefits in different ways (multiple equilibria)

Rents

r high, as politicians can collude, absent separation of powers

Taxes

voters backing government want high τ as they, and their legislators, are residual claimants on revenue

5. Comparative politics results

Presidential vs. parliamentary regimes $\text{no confidence requirement, but separation of powers} \Rightarrow$

- \bullet redistribution more targeted (f to one group)
- less public goods (lower g)
- less rents (lower r)
- smaller size of government (lower τ)

Compare to predictions in other work

Coate-Knight (2011) on US cities find analogous result for τ

D. Summary of predictions

Theoretically predicted effects

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size of government spending
composition into broad vs. targeted programs
(rents extracted by politicians)
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+ or - in Table 2.1

How take these predictions to data? Do predictions hold up? those are the topics of next lecture

Table 2.1
Constitutions and economic policy
Theoretical predictions

Policy outcome	Electoral rules	Form of government
	Majoritarian vs. proportional	Presidential vs. parliamentary
Overall size of government	- /?	_
Composition: broad vs. narrow programs	_	_
Rent extraction	+ / -	-

Empirical Strategy and Results Introduction

How do we test predictions in last lecture as summarized in Table 2.1

Describe data, methods, and selected results

Agenda

- A. Data and their properties
- B. Statistical concerns
- C. Identification: effects on government size
- D. Tests of other predictions

A. Data and their properties

Sample selection

given questions: most interesting variation across countries limit sample to democracies (in retrospect, a mistake!) treat democracy status as random generous definition of democracy, test if matters

Two data sets

average annual obs. in 1990s for 85 democracies (cross section) annual obs. in 1960-98 for 60 democracies (panel)

Policy and performance measures for each question in *Table 2.1* describe in context

Country characteristics

many socioeconomic, historical, geographical, cultural variables covariates with policy or constitutional rules

Measures of electoral rules

binary indicator, MAJ, for electoral formula is only plurality rule used in election of lower house? yes: MAJ = 1, no: MAJ = 0

Measures of forms of government

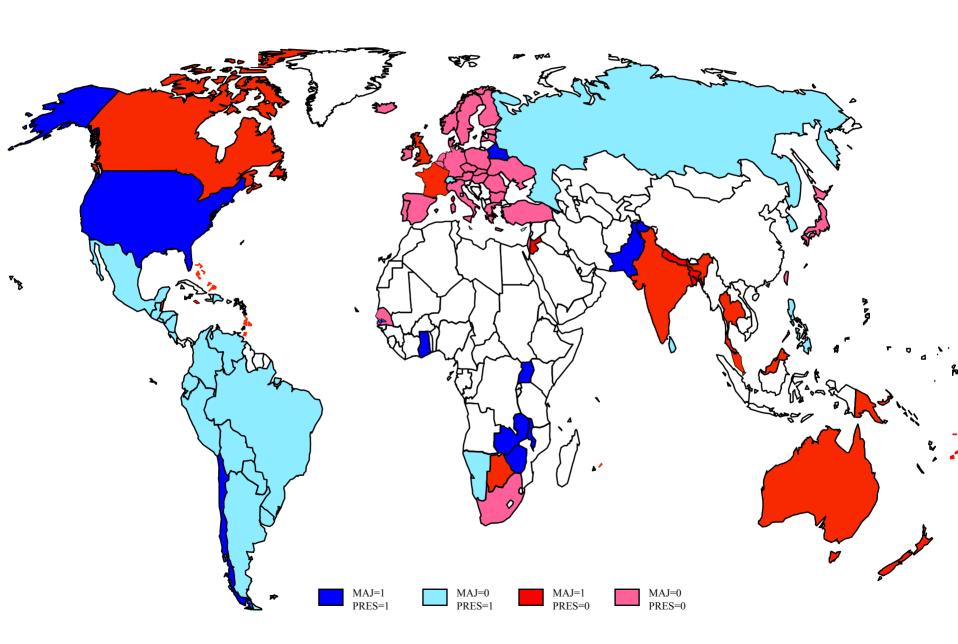
binary indicator, PRES, for confidence requirement is the executive independent of confidence in legislature? yes: PRES = 1, no: PRES = 0 no measure of separation of powers

Main characteristics two features important for empirical strategy

- (i) Constitutional inertia
 deep reforms rare events: panel has no switch in *PRES* five in *MAJ* ten if count mixed systems as well

 and several more in detailed measures of electoral rules
 must estimate constitutional effects from cross-country variation
- (ii) Non-random selection of constitutional rules reflect history, geography and culture *Figure 4.1* must be careful in inference

Constitutional Atlas 1998



B. Statistical concerns

Causal effects or statistical correlations?

serious pitfalls in inference from cross-sectional data
try to address by applying alternative methods

What form of simultaneity

'reverse causation' perhaps not major issue (inertia over 40 years) 'omitted variables' more serious problem

How avoid confounding constitutions and other policy determinants?

hold constant observables which are correlated with policy,
outcomes as well as constitution selection ⇒
use regression with many controls

How about remaining unobservables?

lacking imagination or difficulties in measurement can lead to bias isolate exogenous variation \Rightarrow IV-estimation clean estimates from selection bias \Rightarrow Heckman-style adjustment

How about unwarranted extrapolation for heterogenous groups?

allow for non-linear relations in data and give higher weight to 'local' comparisons ⇒ estimate with matching methods

Illustrate in the context of our problem

use alternative methods to estimate same parameter and discuss specific identifying assumptions in theory and practice apply to constitutional effects on the size of government

C. Identification: effects on government size

Details of empirical strategy for cross-sectional data

what's the parameter of interest?

under what assumptions can we estimate it, given concerns in **B**? what results do we obtain?

Use size of government as specific example

1. Overall question and problem

Parameter of interest

what is direct effect of constitutional reform in country selected at random (ATE)?

How does rule S = 1, vs. S = 0, affect policy outcome Y?

$$\alpha \equiv \mathsf{E}(Y^1 - Y^0) = \mathsf{E}\{E(Y^1 \mid \mathbf{X})\} - \mathsf{E}\{E(Y^0 \mid \mathbf{X})\}$$
 (14)

where last equality relies on law of iterated expectations when E and E refer to unobservables and observables, respectively

More concretely

how does switch PRES or MAJ from 0 to 1 affect central government spending (revenue) as % of GDP?

Problem with observational data

need country-i outcome in two potential states: Y_i^1 , Y_i^0 observe only $Y_i = S_i Y_i^1 + (1 - S_i) Y_i^0$ other outcome is unobserved counterfactual

Pose problem differently

condition on S rather than on \mathbf{X} , rewrite (14) as

$$\alpha = \mathsf{E}(Y^1 - Y^0) = P \cdot [E(Y^1 \mid S = 1) - E(Y^0 \mid S = 1)] \quad (15)$$
$$+ (1 - P) \cdot [E(Y^1 \mid S = 0) - E(Y^0 \mid S = 0)]$$

 α sums effects (ATN and ATT) for countries in states S=0,1 which have probabilities 1-P and P, respectively

(Hypothetical) Experimental data

safe assume e.g.,
$$E(Y^0 \mid S = 1) = E\{E(Y^0 \mid \mathbf{X}, S = 1)\}\$$

= $E\{E(Y^0 \mid \mathbf{X}, S = 0)\} = E(Y^0 \mid S = 0)$

because both unobserved and observed features balanced across S = 0, 1 – by randomization

Yet another way to pose the problem

consider general data-generating process

$$Y_i^S = F^S(\mathbf{X}_i) + \varepsilon_i^S, \quad S = 0, 1$$

$$S_i = \begin{cases} 1 \text{ as } G(\mathbf{W}_i, \mathbf{Z}_i) + \eta_i \ge 0 \\ 0 \text{ otherwise} \end{cases}$$
(16)

where \mathbf{W} is subset of \mathbf{X} , while \mathbf{Z} is not

Identification of α

need strong explicit or implicit assumptions about $F^S(\mathbf{X}), \varepsilon^S, \eta$ to tackle the statistical problems mentioned in \mathbf{B}

2. Linear regression estimates

(i) Recursivity: $Cov(\varepsilon^S, \eta) = 0$

or, equivalently, conditional independence

$$E(Y^1 \mid \mathbf{X}, S = 1) = E(Y^1 \mid \mathbf{X}, S = 0) = E(Y^1 \mid \mathbf{X})$$

selection is random, after controlling for X

- X income, age and quality of democracy, openness, demographics, indicators for federal structure, OECD, continents, colonial history
- (ii) Linearity

$$Y_i^S = F^S(\mathbf{X}_i) + \varepsilon_i^S = \alpha^S + \beta \mathbf{X}_i + \varepsilon_i^S$$

By (i) and (ii),

$$\alpha = \alpha^1 - \alpha^0 = \mathsf{E}\{E(Y^1 - Y^0 \mid \mathbf{X})\}$$

causal effect estimated by coefficient on S in OLS of Y on \mathbf{X}, S

$$Y_i = \alpha^0 + \alpha S_i + \beta \mathbf{X}_i + e_i$$

where
$$e_i = \varepsilon_i^0 + S_i(\varepsilon_i^1 - \varepsilon_i^0)$$

Results in Table 6.1

Table 6.1
Size of government and constitutions
Simple regression estimates

	(1)	(2)	(3)	(4)	(5)	(6)	(7)
Dep. var.	CGEXP	CGEXP	CGEXP	CGREV	CGEXP	CGEXP	CGEXP
PRES	-6.08 (1.97)***	-5.29 (1.92)***		-5.17 (2.44)**	-8.29 (2.72)***	-3.46 (3.88)	-7.49 (2.72)***
MAJ	-3.29 (1.73)*	-5.74 (1.95)***		-3.03 (1.85)	-5.59 (2.68)**	-2.93 (3.09)	-4.81 (2.75)*
PROPRES	,	,	-7.08 (2.70)**	,	,	,	,
MAJPAR			-7.30 (3.02)**				
MAJPRES			-10.36 (2.70)***				
Continents	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Colonies	No	Yes	Yes	Yes	Yes	Yes	Yes
Sample	90s, broad	90s, broad	90s, broad	90s, broad	90s, narrow	60-90s, broad	90s,obs as(6)
Obs.	80	80	80	76	62	60	60
Adj. R2	0.58	0.63	0.63	0.58	0.60	0.54	0.63

Summary

 $PRES = 1 \Rightarrow$ spending more than 5% of GDP smaller $MAJ = 1 \Rightarrow$ slightly smaller effect reflect more rapid growth of government, in 1965-85 constitutional effects appear additive

Is identification convincing?

have we included all the relevant variables in X to rule out 'history and culture' determining both S and Y?

can we trust there are no interaction effects or other non-linearities?

address these in turn in sections 3 and 4

3. Relax conditional independence

Can we rule out selection on unobservables?

i.e., is
$$Cov(e, \eta) \neq 0$$
?

Two prospective sources of selection bias

$$\begin{split} p \lim(\hat{\alpha}) &= \alpha + \mathsf{E}(\varepsilon^1 \mid S = 1) - \mathsf{E}(\varepsilon^0 \mid S = 0) = \\ \alpha + \{\mathsf{E}(\varepsilon^0 \mid S = 1) - \mathsf{E}(\varepsilon^0 \mid S = 0)\} + \mathsf{E}(\varepsilon^1 - \varepsilon^0 \mid S = 1) \end{split}$$

conventional omitted variables: ε^0 and S correlated heterogenous constitutional effects: $\varepsilon^1 - \varepsilon^0$ and S correlated

How relax conditional independence?

- **a.** use instrumental variables to isolate exogenous variation in S
- **b.** adjust estimates, Heckman-style, for correlation e, η

a. Instrumental variables

Identifying assumptions

```
how find \mathbf{Z} such that \mathrm{Cov}(\mathbf{Z},S) \neq 0 and \mathrm{Cov}(\mathbf{Z},e) = 0?
timing of latest constitutional reform (3 indicator variables)
historical waves of reform (hold constant age of democracy)
latitude, fractions speaking English, European language
geographic and cultural 'distance' to old democratic institutions
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Relevant?

constitutional timing: yes, weakly distance measures: yes, definitely

Exogenous?

constitutional timing: yes, a priori distance measures: less certain can test over-identifying restrictions, but low power Apply to our data – results Table 6.2

Compare to earlier estimates in Table 6.1

point estimates agree pretty well note parsimonious 1st stage (weak instruments) standard errors grow as 2nd stage is richer can't reject over-identifying assumptions

Table 6.2
Size of government and constitutions
Heckman and Instrumental Variables estimates

	(0)	(4)
	(3)	(4)
Dep. var.	CGEXP	CGEXP
PRES	-8.65	-4.50
	(3.63)**	(3.89)
MAJ	-3.90	-5.12
•	(3.46)	(3.61)
Conts & Cols	No	COL_UKA,
		LAAM
Sample	90s, broad	90s, broad
Endogenous	PRES	PRES
selection	MAJ	MAJ
Estimation	2SLS	2SLS
rho		
Chi-2: over-id	4.64	3.61
Adj. R2	0.59	0.60
Obs.	75	75

b. Heckman-style adjustment

Well-known idea

estimate selection equation (probit or logit) corresponding to

$$S_i = \begin{cases} 1 & \text{as } G(\mathbf{W}_i, \mathbf{Z}_i) + \eta_i \ge 0 \\ 0 & \text{otherwise} \end{cases}$$

correct estimates of α for remaining correlation $\rho = \text{Corr}(\eta, e)$

Identifying assumptions

same exclusion restrictions on **Z**, as in **a** to avoid relying (only) on functional form

Apply to our data – results Table 6.2

Compare to estimates in Tables 6.1, 6.2

point estimates, if anything, more negative estimated $\rho > 0$ (for PRES) \Rightarrow OLS had positive bias

could allow joint selection of MAJ and PRES, or separate distributions for $\varepsilon^1, \varepsilon^0$ (heterogenous treatment effect)

Table 6.2
Size of government and constitutions
Heckman and Instrumental Variables estimates

	(1)	(2)	(3)	(4)
Dep. var.	CGEXP	CGEXP	CGEXP	CGEXP
PRES	-10.50	-5.37	-8.65	-4.50
	(3.98)***	(2.19)**	(3.63)**	(3.89)
MAJ	-5.69	-4.92	-3.90	-5.12
,	(1.86)***	(2.57)*	(3.46)	(3.61)
Conts & Cols	Yes	Yes	No	COL_UKA, LAAM
Sample	90s, broad	90s, broad	90s, broad	90s, broad
Endogenous	PRES	MAJ	PRES	PRES
selection		,	MAJ	MAJ
Estimation	Heckman	Heckman	2SLS	2SLS
	2-step	2-step		
rho	0.64	-0.02		
Chi-2: over-id			4.64	3.61
Adj. R2			0.59	0.60
Obs.	<i>7</i> 5	75	75	75

4. Relax linearity

Many reasons believe $Y^S = F^S(\mathbf{X})$ non-linear no great concern if \mathbf{X}^1 and \mathbf{X}^0 have similar distribution if not, specification bias – selection on observables – can be severe

Are \mathbf{X}^1 and \mathbf{X}^0 similar?

test $E(\mathbf{X}^1) = E(\mathbf{X}^0)$ suggests not, cf. Table 5.3

PRES: reject in 7 cases out of 9

MAJ: reject in 4 cases out of 9

How take care of prospective specification bias?

parsimonious assumption on functional form rely on 'local' comparisons

Table 5.3
Balancing property
Equal-means tests for different constitutional groups

	Whole	
	sample	
)		
MAJ=1 vs.		
MAJ=0	0.04	
LYP	0.04	
PROP65	0.01	
GASTIL	0.08	
FEDERAL	0.93	
COL_UKA	0.00	
LAAM	0.34	
TRADE	0.44	
PROT80	0.94	
CATHO80	0.00	
PRES=1 vs.		
PRES=1 <i>vs.</i> PRES=0		
LYP	0.00	
	0.00	
PROP65	0.00	
GASTIL	0.00	
FEDERAL	0.22	
COL_UKA	0.44	
LAAM	0.00	
TRADE	0.01	
PROT80	0.03	
CATHO80	0.00	

Probabilities of falsely rejecting the hypothesis of equal means across constitutional groups under the hypothesis of equal variances.

Central idea in matching: mimic experimental measurement split data in 'treated' and 'controls' counterfactual for treated: controls with similar \mathbf{X}_i estimate treatment effect of S on Y non-parametrically

Difficulty

too data-hungry, dimension of X large

Resolution

match on propensity score, rather than directly on \mathbf{X} $p_i = p(\mathbf{X}_i) = \text{Prob}[S = 1 \mid \mathbf{X}_i]$

Identification

- (i) conditional independence of Y given X
- (ii) common-support condition: $0 < p(\mathbf{X}_i) < 1$, all \mathbf{X}_i can rewrite (15) as

$$\alpha = P \cdot \mathsf{E}\{[E(Y^1 \mid p) - E(Y^0 \mid p)] \mid S = 1\} + (1 - P) \cdot \mathsf{E}\{[E(Y^1 \mid 1 - p) - E(Y^0 \mid 1 - p)] \mid S = 0\}$$
 (17)

i.e., condition on p and thus indirectly, not directly, on XNumber of practical questions in evaluating α 1. How estimate $p(\mathbf{X}_i)$?

simple logit (or probit)

specification of \mathbf{X} reflect concern for (trade-off between) conditional-independence and common-support conditions highest t-statistics in equal-means tests and regressions

2. Does matching indeed balance the observations?

Equal-means, for same X as before, in three strata for p PRES: now reject only in 2 out of 27 cases (cf. Table~5.3)

MAJ: again, reject in 2 out 27 cases

3. How estimate P?

relative sample frequency of S = 1

Table 5.3

Balancing property

Equal-means tests for different constitutional groups

	Whole sample	p < 0.33	0.33 < p < 0.67	0.67 < p
	Sample			
MAI=1 vs.				
MAJ=0				
LYP	0.04	0.04	0.62	0.21
PROP65	0.01	0.32	0.90	0.04
GASTIL	0.08	0.33	0.55	0.37
FEDERAL	0.93	0.79	0.57	0.48
COL_UKA	0.00	0.69	0.42	0.35
LAAM	0.34	0.27	0.39	0.17
TRADE	0.44	0.13	0.93	0.31
PROT80	0.94	0.56	0.75	0.37
CATHO80	0.00	0.11	0.46	0.83
PRES=1 vs.				
PRES=0				
LYP	0.00	0.87	0.01	0.54
PROP65	0.00	0.34	0.39	0.86
GASTIL	0.00	0.59	0.22	0.71
FEDERAL	0.22	0.07	0.30	0.27
COL_UKA	0.44	0.88	0.56	0.83
 LAAM	0.00	0.53	0.23	0.22
TRADE	0.01	0.33	0.34	0.40
PROT80	0.03	0.65	0.60	0.22
CATHO80	0.00	0.28	0.24	0.02

Probabilities of falsely rejecting the hypothesis of equal means across constitutional groups under the hypothesis of equal variances. Strata defined on the common support of propensity scores, p, estimated by logit regressions including: *LYP*, *PROP65*, *GASTIL*, *FEDERAL*, *COL_UKA*, *LAAM*.

- 4. How estimate $E\{E(Y^1 \mid p) \mid S = 1\}$? sample mean among "treated"
- 5. How estimate $\mathsf{E}\{E(Y^0 \mid p) \mid S = 1\}$? which controls matched with given p_i among treated?
 - (i) nearest-neighbor: the (one) control with closest p_i produces (mostly) natural matches, cf. Table 5.2
 - (ii) stratification: arithmetic mean, all controls in same p interval (iii) kernel: geometric mean, all controls in radius of p_i
- 6. How impose common-support condition (comparability)? compute 3-5 only for overlapping support of p_i

Table 5.2 Estimated propensity scores

(a) Majoritarian elections

Country	PSCORE	MAJ	Country	PSCORE	MAJ
Uruguay	0.052	0	Nepal	0.337	1
Sweden	0.070	0	South Korea	0.355	0
Greece	0.073	0	Bangladesh	0.371	1
Bulgaria	0.075	0	Philippines	0.377	1
Italy	0.077	0	Namibia	0.419	0
UK	0.078	1	Barbados	0.496	1
Romania	0.083	0	New Zeland	0.568	1
Peru	0.084	0	Jamaica	0.582	1
Belgium	0.090	0	Ireland	0.617	0
Norway	0.090	0	Canada	0.641	1
France	0.093	1	Singapore	0.659	1
Spain	0.095	0	Israel	0.673	0
Latvia	0.093	0	Sri Lanka	0.674	0
Portugal	0.101	0	Trinidad&Tobago	0.674	1
Portugai Denmark	0.104	0	Australia	0.694	1
	0.105		South Africa	0.755	
Hungary		0			0
Japan Colombia	0.108	1	Cyprus (G) Malta	0.759	0
	0.112	0		0.760	0
Estonia	0.114	0	Bahamas	0.763	1
Guatemala	0.115	0	Pakistan	0.781	1
Czech Republic	0.126	0	Uganda	0.790	1
Luxembourg	0.127	0	Gambia	0.794	1
Chile	0.128	1	Ghana	0.797	1
Argentina	0.132	0	Zimbabwe	0.808	1
Finland	0.132	0	Belize	0.812	1
Paraguay	0.133	0	Fiji	0.828	0
Slovak Republic	0.141	0	Malawi	0.831	1
Nicaragua	0.148	0	St. Vincent&Granada	0.856	1
Dominican Republic	0.152	0	Zambia	0.856	1
Netherlands	0.153	0	Malaysia	0.857	1
Ecuador	0.157	0	Mauritius	0.873	1
Germany	0.160	0	India	0.886	1
Russia	0.161	0	Papua New Guina	0.904	1
Poland	0.177	0	Botswana	0.924	1
Bolivia	0.181	0			
Honduras	0.185	0			
Mexico	0.194	0			
Austria	0.199	0			
Iceland	0.212	0			
Switzerland	0.214	0			
Turkey	0.220	0			
Brazil	0.230	0			
Costa Rica	0.240	0			
El Salvador	0.258	0			
Thailand	0.264	1			
Venezuela	0.292	0			
USA	0.297	1			
Senegal	0.320	0			

PSCORE is the predicted value of a logit regression of *MAJ* on *LYP*, *PROP65*, *FEDERAL*, *GASTIL*, *LAAM*, *COL_UKA*

Boldface observations are discarded to impose common support.

Apply to our data – results Table 6.3

Compare to earlier estimates in Tables 6.1 - 6.2

point estimates agree

based on fewer observations on common support

higher standard errors: trade off less bias against less efficiency

Summary: estimated constitutional effects

consistent with theory

larger government, by $\sim 5\%$ of GDP, of parliamentary democracy by about same amount of proportional democracy

Table 6.3
Size of government and constitutions
Matching estimates

	(1)	(2)	(3)	(4)	(5)	(6)
Dep. var.	CGEXP	CGEXP	CGEXP	CGEXP	CGEXP	CGEXP
PRES MAJ	-7.30 (2.30)*** -5.76	-7.91 (2.90)*** -6.55	-5.87 (4.93) -4.87	-7.92 (5.11) -4.08	-2.54 (2.30) -6.59	-4.00 (3.45) -8.81
2722	(2.94)*	(2.82)**	(3.65)	(4.16)	(3.06)**	(3.15)***
Estimation Sample Logit Specif.	Kernel 90s, broad 1	Kernel 90s, broad 2	Strat 90s, broad 1	Strat 90s, broad 2	Nearest 90s, broad 1	Nearest 90s, broad 2
Obs. on common support	65 <i>PRES</i> 67 <i>MAJ</i>	40 <i>PRES</i> 57 <i>MAJ</i>	65 PRES 67 MAJ	40 <i>PRES</i> 57 <i>MAJ</i>	65 PRES 67 MAJ	40 <i>PRES</i> 57 <i>MAJ</i>

5. Summary: size of government

Estimated constitutional effects

consistent with theory

larger government, by $\sim 5\%$ of GDP, of parliamentary democracy by about same amount of proportional democracy

D. Tests of other predictions

Quick description of results

1. Composition of government

Measurement

Y welfare-state spending (social transfers), as % of GDP

X same as for overall spending

Estimation

same battery of methods as in C

Summary of findings

partly consistent with theory

2-3% of GDP higher in (good, old) proportional democracies

2-3% of GDP higher in (good, old) parliamentary democracies

2. Political rents and corruption

Measurement

Y perception indexes for corruption, and government (in)effectiveness, 0-10 scale

X dozen covariates with corruption, suggested by earlier studies

Here can also exploit (piecemeal) electoral reforms

district magnitude & ballot structure, as well as electoral formula cross-section and panel (fixed-effect) estimates agree

Summary of findings: electoral rules

larger electoral districts: less corruption (and inefficiency) more list voting: more corruption (and inefficiency) quantitatively important effects binary indicator: no robust effect

Summary of findings: form of government less corruption in presidential regimes? perhaps, but only in better democracies

3. Extensions – examples of more recent research

- (i) Deeper theoretical and empirical analysis mechanism for higher government spending under PR? incentives for politicians vs. indirect effect via party system
- (ii) Wider scope of empirical work systematic effects beyond fiscal policy (and corruption)? trade policy, regulatory policy, economic performance
- (iii) More extensive data sets with sharper identification from time variation around reforms? exploit switches in and out of democracy